

# Performance of Cramer Rao Bound on Jitter Variance Estimation for Blind Channel

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**Abstract:** In this paper, we compare the Cramer-Rao bounds (CRB) under the linear and norm constraints, for symbol and channel response estimation. We consider the problem of estimating parameters of an irregular sampling process defined as a uniform sampling process in which the deviations from the nominal sampling times constitute a random independent identically distributed (IID) process (jitter). For multi channel FIR systems (SIMO), it has been shown that the norm-constrained CRB is smaller than the linear-constrained CRB. And also the estimator's variance is compared to the linear and norm constraints. Computer simulations verify the effectiveness of the proposed estimation method for small jitter.

**Keywords:** jitter, cramer rao bounds, vector, SNR,

## 1 . INTRODUCTION

In modern communication, SNR estimation ply more important role. First in this process, analog signal converts into discrete signal and this sampling process consist of a systematic random timing error (jitter). A related problem to jitter detection is the problem of estimation of jitter parameters. An efficient estimation procedure can be applied, for example, in classification of jammers and identification of friendly sources in ECM systems. Unlike a jitter detection, which under the conditions stated above, can be done only in the bi-spectral domain. It is provided that the sampling rate is no smaller than Nyquist rate, estimation of jitter parameters can be done also the spectral domain. The question as to which domain is preferable for carrying out the estimation is considered and it is shown using Cramer-Rao bounds. The answer depends on the skewness of the continuous process. If the skewness is larger than one, then its noise variance decreases. In this work we present an algorithm for estimation of the jitter variance using the bi-spectrum. This procedure can be adapted to the estimation of other jitter parameters, as long as they affect by the characteristic function of the jitter. The advantage of estimating jitter variance is that the jitter detection and estimation problems can be solved simultaneously. Specifically, the sampling is assumed to be uniform with additive independent identically distributed (IID) jitter.

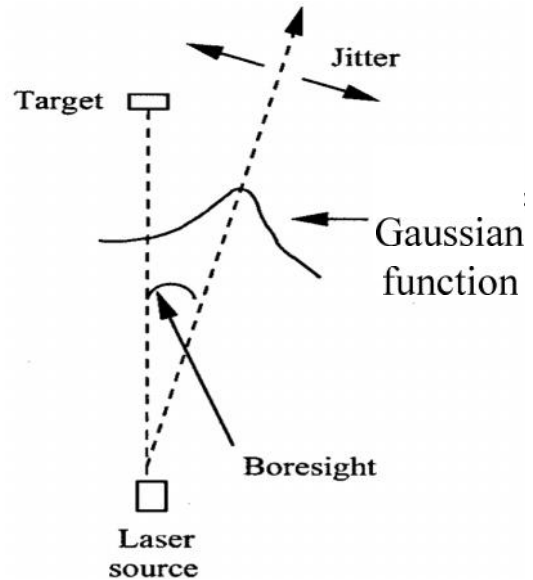


Fig1: jitter

A key question here: what is the ultimate performance limit that any estimation technique can achieve with respect to jitter and boresight estimation using the return photon signal? In this paper, we derive the Cramer–Rao lower bound (CRLB) on this estimator variance. And also compare the performance of linear and norm constrains by using CRLB. With the help of Gaussian far-field irradiance profile, we consider  $N$  shots directed from a laser source at a target in the far field as shown in Fig. 1.

IID process is a discrete time random process,  $X(n)$ , (or random sequence  $X(1), X(2) \dots$ ) is called an IID process.  $E[X(n)] = m$  and  $E[X^2(n)] = m^2 + \sigma^2$  if  $n_1 = n_2$  and  $E[X(n_1)X(n_2)] = m^2$  if  $n_1 \neq n_2$ .

$$R_X(n_1, n_2) = E[X(n_1)X(n_2)] \tag{1}$$

$$= \begin{cases} E[X(n_1)] \cdot E[X(n_2)], & \text{if } n_1 \neq n_2 \\ E[X^2(n_1)], & n_1 = n_2 \end{cases}$$

$$= \begin{cases} m^2 & , \text{if } n_1 \neq n_2 \\ m^2 + \sigma^2 & , n_1 = n_2 \end{cases}$$

where

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$$\delta_{n_1, n_2} = \begin{cases} 1, & n_1 = n_2 \\ 0, & n_1 \neq n_2 \end{cases}$$

$$\delta_{n_1, n_2} = \delta(n_1 - n_2)$$

$$C_X(n_1, n_2) = R_x(n_1, n_2) - m_x(n_1)m_x(n_2)$$

$$C_X(n_1, n_2) = \sigma^2 \cdot \delta_{n_1, n_2}$$

First let us consider a communication system consisting of  $M$  channels. The complex finite impulse response of the  $i^{th}$  channel is denoted by  $h_i(0), \dots, h_i(L)$ ,  $a = 1, 2, \dots, h_i(L) \neq 0$  for at least one  $I$  and  $h_j(0) \neq 0$  for at least one  $j$ . All channels are driven by the same complex input symbol sequence so...,  $S_{N+L} \sim$  transmitted from a single user. If we let  $\mathbf{y}_i(\mathbf{n}), \mathbf{n} = 0, \dots, N-1$ , be the output samples at  $N$  different time instants, for the  $i^{th}$  channel, the output samples of the  $i^{th}$  channel can be written by (2) and its parameters written as (3), (4), (5) & (6). Jitter variance represented by (9)

$$\mathbf{y}_i = \mathbf{H}_i \mathbf{s} + \mathbf{w}_i = \tilde{\mathbf{S}} \mathbf{h}_i + \mathbf{w}_i, \quad i = 1, 2, \dots, M \tag{2}$$

where  $\mathbf{y}_i = [y_i(0), y_i(1), \dots, y_i(N-1)]^T$ ,

$$\mathbf{H}_i = \begin{bmatrix} h_i(L) & \dots & h_i(0) & & \\ & \ddots & & \ddots & \\ & & h_i(L) & \dots & h_i(0) \\ & & & \ddots & \\ & & & & h_i(L) & \dots & h_i(0) \end{bmatrix}_{N \times (N+L)} \tag{3}$$

$$\mathbf{s} = [s_0, \dots, s_{N+L-1}]^T \tag{4}$$

$$\tilde{\mathbf{S}} = \begin{bmatrix} s_L & \dots & s_0 \\ s_{L+1} & \dots & s_1 \\ \vdots & \ddots & \vdots \\ s_{N+L-1} & \dots & s_{N-1} \end{bmatrix} \tag{5}$$

$$\mathbf{h}_i = [h_i(0), \dots, h_i(L)]^T \tag{6}$$

Blind symbol/channel estimation is to estimate symbol vector  $\mathbf{s}$  and channel response vector  $\mathbf{h} = [h_1, \dots, h_M]$ , from the noisy measurement  $\mathbf{y}_i, i = 1, 2, \dots, M$ . As is shown in [4, 10], one complex parameter in  $[\mathbf{s}^T, \mathbf{h}^T]$  is inherently non-identifiable. To obtain a unique solution, constraints have to be incorporated into symbol/channel estimation approaches. Two widely used constraints include (a) the linear constraint: the  $l$ -th (complex) element of the solution to symbol (of channel response) vector is set to one, and (b) the norm constraint: the norm of the solution to symbol (or channel response) vector is set to one and the  $l$ -th element of the concerned solution vector is restricted to be real

**Problem Formulation**

Let us assume the noise component  $\mathbf{W}_i$  is independent, jitter variance  $\sigma_w^2$  and  $\mathbf{w}_i, i = 1, 2, \dots, M$ , are also independent of each other. Denote the linear-constrained symbol vector by

(7) and the norm-constrained symbol vector by (8). For real symbols/channel response, linear and norm constraints were discussed in (3)-(5). In the context of complex symbols, the linear constraint functions are given by equation

$$\mathbf{f}_l = \begin{Bmatrix} \Re e(\gamma_{l_0}) - 1 \\ \Im m(\gamma_{l_0}) \end{Bmatrix} = \mathbf{0} \tag{7}$$

$$\mathbf{f}_n = \begin{Bmatrix} (||\rho||^2 - 1)/2 \\ \Im m(\rho_{l_0}) \end{Bmatrix} = \mathbf{0} \tag{8}$$

$$\frac{\partial \mathbf{f}_l}{\partial \theta_l^T} = [\mathbf{0}_{2 \times 2M(L+1)}, \mathbf{G}_l^T]$$

$$\frac{\partial \mathbf{f}_n}{\partial \theta_n^T} = [\mathbf{0}_{2 \times 2M(L+1)}, \mathbf{G}_n^T]$$

Where,

$$\mathbf{G}_l = [\mathbf{e}_{l_0}, \mathbf{e}_{N+L+l_0}]$$

$$\mathbf{G}_n = \left[ \begin{bmatrix} \Re e(\rho) \\ \Im m(\rho) \end{bmatrix}, \mathbf{e}_{N+L+l_0} \right]$$

$$\hat{\mathbf{S}} = \begin{bmatrix} \Re e(\mathbf{S}) & -\Im m(\mathbf{S}) \\ \Im m(\mathbf{S}) & \Re e(\mathbf{S}) \end{bmatrix}$$

$$\hat{\mathbf{H}} = \begin{bmatrix} \Re e(\mathbf{H}) & -\Im m(\mathbf{H}) \\ \Im m(\mathbf{H}) & \Re e(\mathbf{H}) \end{bmatrix}$$

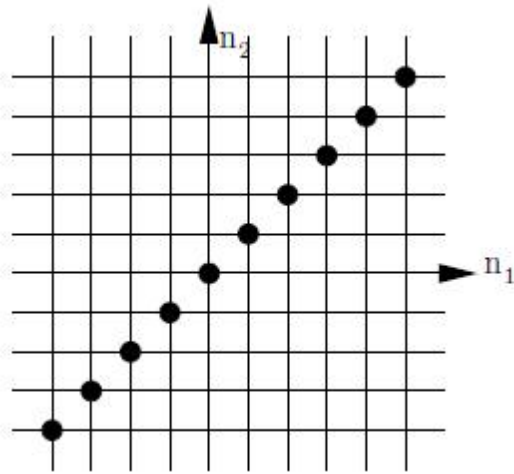
$$K = \begin{bmatrix} \sigma_w^2 & 0 & 0 \\ 0 & \sigma_w^2 & 0 \\ 0 & 0 & \sigma_w^2 \end{bmatrix} \tag{9}$$

Assume that mean IID of both noise and signals are zero with variances  $\sigma_{2n}$  and  $\sigma_{2s}$ , respectively, the covariance of the channel estimation error is approximated by the absolute CRB should be obtained from the most original equation (9). However, since the performance of the provided algorithm is only related (12), we would like to resort to the CRB after the pre-processing of our re-modulation. By using this CRB, we can also build up its relationship with the MSE that is derived in the previous subsection. The deterministic CRB for CPSOS will be considered. Here, where the observations are  $\mathbf{z} = [\mathbf{z}^T 1, \dots, \mathbf{z}^T M]^T$ , and the parameters to be estimated are  $\theta = [\text{vec}(\mathbf{H}), \mathbf{d}_1, \dots, \mathbf{d}_M, \sigma_{2n}]$ .

To calculate CRB, we need the joint probability density function (PDF) of  $\mathbf{z}$ , denoted as  $p(\mathbf{z}|\theta)$ . Since  $\eta_i$  is correlated with both  $\eta_{i-1}$  and  $\eta_{i+1}$ , the covariance matrix of  $\mathbf{z}$ , denoted as  $\mathbf{R}_z$ , is an  $MJ(N+L) \times MJ(N+L)$  Toeplitz matrix with the main diagonal elements 2, the  $(JN+1)$ th,  $-(JN+1)$ th diagonal elements  $-1$ , and all other elements are 0. We note that the inverse of such a huge Toeplitz matrix is mathematically represented as (10,11).

$$ACRB_{vec(H)} = I_K \otimes \left( \frac{\sigma_w^2 (K^H)^T K^T}{2M\sigma_w^2} \right) \tag{10}$$

$$ACRB_{vec(H(1,k))} = \frac{\sigma_w^2 (K^H)^T K^T}{2M\sigma_w^2} \tag{11}$$



$$r[n] = K \exp\left\{-\frac{[(x[n] + A)^2 + y^2[n]]}{2\omega^2}\right\}, \quad n = 1, 2, \dots, N, \tag{12}$$

where  $K$  is a gain constant,  $\omega$  is the noise standard deviation of the far-field irradiance pattern along any direction,  $x[n]$  and  $y[n]$  are the angular coordinates for the  $n^{\text{th}}$  observation, and  $A$  is the boresight. The boresight represents the mean offset between the beam peak and the desired target (Fig. 1). The exact beam pointing is a random variable due to factors such as tracking errors, mechanical vibrations, and atmo where  $\sigma_j$  is the jitter standard deviation along any direction. In our subsequent discussions, we refer to  $\sigma_j$  simply as jitter. The target is assumed to be a point target in our model. The log pdf of the observation vector  $\mathbf{r} = [r[1], r[2], \dots, r[N]]$  for a given set of  $K, A$ , and  $\sigma_j$  is obtained by following equations.

$$\begin{aligned} \log p(\mathbf{r}; A, \sigma_j) &= 2N \log \Omega - \sum_{n=1}^N \log r[n] - 2N \log \sigma_j \\ &\quad - \frac{1}{2\sigma_j^2} \left[ A^2 N + 2\Omega^2 \sum_{n=1}^N \log(K / r[n]) \right] \\ &\quad - \sum_{n=1}^N \log \left\{ I_0 \left[ \frac{A}{\sigma_j^2} \sqrt{2\Omega^2 \log(K / r[n])} \right] \right\} \end{aligned}$$

$$\frac{\partial \log p(\mathbf{r}; A, \sigma_j)}{\partial A} = -\frac{AN}{\sigma_j^2} + \frac{1}{\sigma_j^2} \sum_{n=1}^N \frac{I_1(u[n])}{I_0(u[n])} \sqrt{z[n]}, \tag{13}$$

$$\begin{aligned} \frac{\partial \log p(\mathbf{r}; A, \sigma_j)}{\partial \sigma_j} &= -\frac{2N}{\sigma_j} + \frac{A^2 N}{\sigma_j^3} + \frac{1}{\sigma_j^3} \sum_{n=1}^N z[n] \\ &\quad - \frac{2}{\sigma_j} \sum_{n=1}^N \frac{I_1(u[n])}{I_0(u[n])} u[n] \end{aligned} \tag{14}$$

$$\begin{aligned} J_{11} &= \frac{N}{\sigma_j^2} - \frac{N}{\sigma_j^4} \times \\ &\quad \frac{1}{2\pi\sigma_j^2} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} S_{11} \times \exp\left(-\frac{x^2[n] + y^2[n]}{2\sigma_j^2}\right) dx[n] dy[n] \end{aligned} \tag{15}$$

$$\begin{aligned} j_{12} = j_{21} &= -\frac{2AN}{\sigma_j^3} - \frac{2AN}{\sigma_j^5} \times \\ &\quad \frac{1}{2\pi\sigma_j^2} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} S_{12} \times \exp\left(-\frac{x^2[n] + y^2[n]}{2\sigma_j^2}\right) dx[n] dy[n] \end{aligned} \tag{16}$$

$$\begin{aligned} j_{22} &= -\frac{2N}{\sigma_j^2} + \frac{3A^2 N}{\sigma_j^4} + \frac{N}{\sigma_j^4} \times \\ &\quad \frac{1}{2\pi\sigma_j^2} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} S_{22} \times \exp\left(-\frac{x^2[n] + y^2[n]}{2\sigma_j^2}\right) dx[n] dy[n] \end{aligned} \tag{17}$$

Where var denotes variance and  $\mathbf{J}^{-1}_{mn}$  denotes the  $m, n^{\text{th}}$  element of inverse matrix  $\mathbf{J}^{-1}$ . Thus, inequalities (16,17) provide the performance limits when the boresight and jitter are jointly estimated. They can be numerically evaluated, preferably in the log domain for better numerical behaviour. If only boresight is estimated for a known jitter, we get a bound  $\text{var} \hat{A} \geq 1/J_{11}$ . Similarly, when the boresight is known, the bound on jitter estimation variance is  $\text{var} \hat{\sigma}_j \geq 1/J_{22}$ . The bounds for joint parameter estimation are higher than individual parameter estimation. In the special case of zero boresight, a closedform expression for CRLB on jitter estimate can be obtained by substituting  $A=0$  into Eq. (13,14). This is given by

$$\text{var}(\hat{\sigma}_j) \geq \sigma_j^2 / 4N \tag{18}$$

WHERE

$$P(x[n], y[n]) = \frac{1}{2\pi\sigma_j^2} \exp\left[-\frac{(x^2[n] + y^2[n])}{2\sigma_j^2}\right]$$

$$CRB(\|\gamma\|) = \frac{\sigma_{ip}^2}{2} \text{tr} \left[ (V_l^T H^T P \frac{1}{S} H V_l)^{-1} \right]$$

$$CRB(\|\rho\|) = \frac{\sigma_{\omega}^2}{2} \text{tr} \left[ (V_n^T H^T P \frac{1}{S} H V_n)^{-1} \right] / \|\gamma\|^2$$

$$CRB(\|\gamma\|) / \|\gamma\|^2 > CRB(\|\rho\|)$$

$$H^T P \frac{1}{S} H \begin{bmatrix} RE(\rho) \\ IM(\rho) \end{bmatrix} = 0$$

$$J_{\eta} = \frac{2}{\sigma_{\omega,C}^2} B^T B$$

WHERE

$$B = \begin{bmatrix} \text{Re}(Q) & \text{Im}(Q) & \text{Re}(F) & \text{Im}(F) \\ -\text{Im}(Q) & \text{Re}(Q) & -\text{Im}(F) & \text{Re}(F) \end{bmatrix}$$

$$Q = I_N \otimes [C_1 h_1, \dots, C_p h_p]$$

$$F = [F_1^T, \dots, F_N^T]^T$$

$$F_N = [s_1(n)C_1, \dots, s_p(n)C_p]$$

$$CRB(\|\gamma_C\|) = \frac{\sigma_{\omega,C}^2}{2} \text{tr} \left[ (V_{l,c}^T Q^T P \frac{1}{F} Q V_{l,c})^{-1} \right]$$

$$CRB(\|\rho_C\|) = \frac{\sigma_{\omega,C}^2}{2} \text{tr} \left[ \left( V_{n,c}^T Q^T P \frac{1}{F} Q V_{n,c} \right)^{-1} \right] / \|\gamma_c\|^2$$

$$\eta^T = \begin{bmatrix} \text{Re}(s_1)^T, \dots, \text{Re}(s_N)^T, \text{Im}(s_1)^T, \dots, \text{Im}(s_N)^T, \text{Re}(h_1)^T \\ \dots, \text{Re}(h_p)^T, \text{Im}(h_1)^T, \dots, \text{Im}(h_p)^T \end{bmatrix}$$

Figure 1 shows the two CRBs and experimental mean square errors for constrained symbol estimation, versus SNR where SNR = 20 log<sub>10</sub>(I/hlluk/u.) (dB). The mean squared errors of the constrained channel response estimation and the corresponding CRBs versus SNR are shown in Figure 2. The mean squared error of the constrained channel response estimation is defined in the same way as that of the constrained symbol estimation. It can be seen that the relative statistical efficiency under the two constraints remain the same over a wide range of SNR. For symbol estimation, at SNR below 30 dB, the performance breaks away from the CRBs. For channel response estimation, it is interesting to note that the performance becomes better than th& CRBs at SNR below 20 dB, and the difference can be as large as around 4 dB at SNR = 'U dB. Such a phenomenon was reported in [4] and recently in [8]. As explained in [11], this peculiar observation takes place because an upper bound on the variance of a norm-constrained error. goes below CRBs.

$$U_0^H R_\omega^{-1/2} C_n H = 0, \quad n = 1, \dots, N \tag{19}$$

where C<sub>n</sub> is the J(N+L) × J(L+1) Toeplitz matrix with the first column e(n-1)J+1 and ep is defined as the pth column of LJ(N+L). The first row of C<sub>n</sub> is [1, 01 × (J(L+1)-1)] for n = 1 and is 01 × J(L+1) for n ≥ 2.

Define

$$K = \begin{bmatrix} C_1^H R_\omega^{-1/2} U_0, C_2^H R_\omega^{-1/2} U_0, \dots, C_N^H R_\omega^{-1/2} U_0 \end{bmatrix} \tag{20}$$

The channel matrix H could be estimated from K<sup>H</sup>H = 0. Therefore, the estimate of H, denoted as H-hat, is a basis matrix of the orthogonal complement space of K. We will show later that the dimension of the orthogonal complement space of K is exactly K. Therefore, H-hat can be obtained from left singular vectors of K and is away from the true H by an unknown matrix B, mathematically represented by following equations

$$H-hat = HB \tag{21}$$

$$G = G-hat(I_N \otimes B^{-1}) \tag{22}$$

$$Z_i = G-hat(I_N \otimes B^{-1})d_i + \eta_i \tag{23}$$

$$\tilde{x}_i(n) = \begin{bmatrix} \tilde{x}_i^{(1)}(n), \tilde{x}_i^{(2)}(n), \dots, \tilde{x}_i^{(j)}(n) \end{bmatrix}^T \tag{24}$$

$$s_i(n) = \begin{bmatrix} s_i^{(1)}(n), s_i^{(2)}(n), \dots, s_i^{(k)}(n) \end{bmatrix}^T \tag{25}$$

$$\omega_i(n) = [\omega_i^{(1)}(n), \omega_i^{(2)}(n), \dots, \omega_i^{(j)}(n)]^T \tag{26}$$

$$\tilde{H}(n) = \begin{bmatrix} \tilde{h}^{(1,1)}(n) & \dots & \tilde{h}^{(1,K)}(n) \\ \vdots & \ddots & \vdots \\ \tilde{h}^{(J,1)}(n) & \dots & \tilde{h}^{(J,K)}(n) \end{bmatrix} \tag{27}$$

$$\tilde{X}_i(n) = \tilde{H}(n)s_i(n) + \omega_i(n) \tag{28}$$

II. RESULTS:

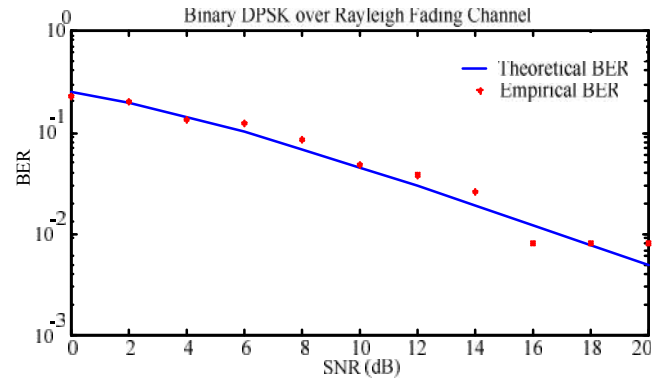
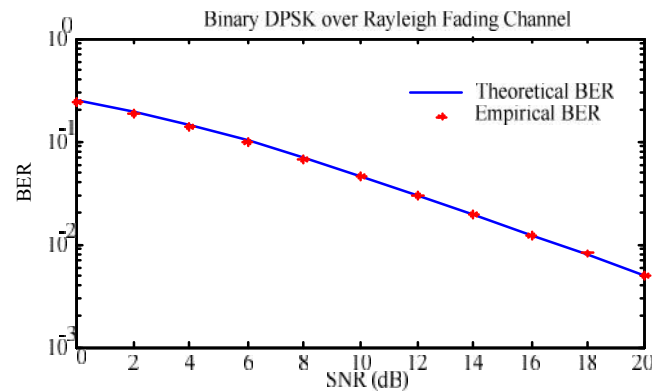


Fig :3 SNR estimation linear constraints



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Fig :4 norm constraints

From figure 3, it is clearly observed that Cramer-Rao bounds under linear constraint method the estimated SNR value could not follow the theoretical results. But figure 4 portrays that Cramer-Rao bounds under Norm constraints method the estimated SNR value follows the theoretical results. So that the norm constraints method is suitable for SNR estimation for blind channel.

Conclusion

In this paper it shows that for random process (jitter variance) is applied to the linear and norm constraints the SNR was calculated and it shows good performance given as graph representation like norm constraints greater than linear constraints it is suggested that it can able to apply for SNR estimation to the blind channel

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